Lecture 2

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Outline

Concentration Inequalities Revisited

Universal Family of Hash Functions

Counting Distinct Items Analysis of Algorithm from Lecture 0

AMS Algorithm for Counting Distinct Element

First and Second Moment Bounds

Markov Inequality For any positive random variable X and t > 0
F[X]

$$\Pr[X > t] \le \frac{\mathsf{E}[X]}{t}$$

Chebyshev Inequality For any random variable X and t > 0

$$\Pr[|X - E[X]| > t] \le \frac{\operatorname{Var}[X]}{t^2}$$

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The Chernoff Bound

• Let $X_1, X_2...X_n$ be *n* independent Bernoulli random variables with $\Pr(X_i = 1) = p_i$. Let $X = \sum X_i$. Hence, $E[X] = E\left[\sum X_i\right] = \sum E[X_i] = \sum \Pr(X_i = 1) = \sum p_i = \mu(say)$

Then the Chernoff Bound says for any $\epsilon > 0$

$$\mathsf{Pr}(X > (1 + \epsilon)\mu) \leq \left(rac{e^{\epsilon}}{(1 + \epsilon)^{\epsilon}}
ight)^{\mu}$$
 and $\mathsf{Pr}(X < (1 - \epsilon)\mu) \leq \left(rac{e^{-\epsilon}}{(1 - \epsilon)^{1 - \epsilon}}
ight)^{\mu}$

When $0 < \epsilon < 1$ the above expression can be further simplified to

$$\Pr(X > (1 + \epsilon)\mu) \le e^{-\frac{\mu\epsilon^2}{3}}$$
and $\Pr(X < (1 - \epsilon)\mu) \le e^{-\frac{\mu\epsilon^2}{2}}$

Hence

$$\Pr(|X-\mu| > \epsilon\mu) \le 2e^{rac{-\mu\epsilon^2}{(-3)}}$$
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Universal Hash Family

A family of hash functions $\mathcal{H} = \{h \mid h : [N] - - > [M]\}$ is called a pairwise independent family of hash functions if for all $i \neq j \in [N]$ and any $k, l \in [M]$

$$\Pr_{h\leftarrow\mathcal{H}}[h(i)=k\wedge h(j)=l]=rac{1}{M^2}$$
strongly universal hash family (1)

Hash functions are uniform over [M],

$$\Pr_{h \leftarrow \mathcal{H}} [h(i) = k] = \frac{1}{M}$$
(2)

$$\Pr_{h \leftarrow \mathcal{H}}[h(i) = h(j)] = \frac{1}{M}$$
weakly universal hash family (3)

► Construction Let *p* be a prime. For any

$$a, b \in \mathbb{Z}_p = \{0, 1, 2, ..., p - 1\}$$
, define $h_{a,b} : \mathbb{Z}_p \to \mathbb{Z}_p$ by
 $h_{a,b}(x) = ax + b \mod p$. Then the collection of functions
 $\mathcal{H} = \{h_{a,b} | a, b \in \mathbb{Z}_p\}$ is a pairwise independent hash family.

Algorithm 1 $[a, \epsilon, \delta]$

$$\begin{split} & \epsilon' = \epsilon/2 \\ & \text{for } t = 1, \lceil (1 + \epsilon') \rceil, \lceil (1 + \epsilon')^2 \rceil, ... \lceil (1 + \epsilon')^{\log_{1 + \epsilon'} n} \rceil \text{ do} \\ & \delta' = \frac{\epsilon' \delta}{\log n} \{ \text{Run in parallel} \} \\ & b_t = \text{ESTIMATE}(\mathbf{a}, t, \epsilon', \delta') \{ b_t \text{ is a boolean variable YES}/ \\ & \text{NO} \} \\ & \text{end for} \\ & \text{return the smallest value of } t \text{ such that } b_{t-1} = \text{YES and } b_t = \text{NO}, \end{split}$$

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if no such t exists, return n

Algorithm 2 [ESTIMATE($\mathbf{a}, t, \epsilon', \delta'$)]

 $count \leftarrow 0$

for
$$i=1$$
 to $rac{c}{\epsilon'^2}\lograc{1}{\delta'}$ do

Select a hash function h_i uniformly and randomly from a fullyindependent hash family \mathcal{H} {run in parallel}

$$b_t^i \leftarrow \mathsf{NO}$$

repeat

Consider the current element in the stream **a**, say $a_l = (j, \nu)$ if $h_i(j) == 1$ then $b_t^i \leftarrow YES$, BREAK end if until **a** is exhausted if $b_t^i == NO$ then count = count + 1end if end for

Algorithm 3 [ESTIMATE($\mathbf{a}, t, \epsilon', \delta'$)]continued

 $\begin{array}{l} \text{if } count \geq \frac{1}{e} \frac{c}{\epsilon'^2} \log \frac{1}{\delta'} \text{ then} \\ \text{return NO} \\ \text{else} \\ \text{return YES} \\ \text{end if} \end{array}$

- Space Complexity: $O(\frac{1}{\epsilon^3} \log n(\log \frac{1}{\delta} + \log \log n + \log \frac{1}{\epsilon}))$
- Time Complexity: $O(\frac{1}{\epsilon^3} \log n(\log \frac{1}{\delta} + \log \log n + \log \frac{1}{\epsilon}))$

Lemma

Consider the ith round of ESTIMATE($\mathbf{a}, t, \epsilon', \delta'$) for any $i \in [\frac{C}{\epsilon^2} \log \frac{1}{\delta'}]$

▶ If
$$DE > (1 + \epsilon')t$$
 then $\Pr[b_t^i == NO] \le \frac{1}{e} - \frac{\epsilon}{2e}$.
▶ If $DE < (1 - \epsilon')t$ then $\Pr[b_t^i == NO] \ge \frac{1}{e} + \frac{\epsilon}{2e}$.

Lemma

▶ If
$$DE > (1 + \epsilon')t$$
 then $\Pr[b_t == NO] \le \frac{\delta'}{2}$.
▶ If $DE < (1 - \epsilon')t$ then $\Pr[b_t == YES] \le \frac{\delta'}{2}$.

Lemma

• If
$$|DE - t| > \epsilon' t$$
 then $\Pr[ERROR] \le \delta'$.

Lemma

For all t such that $|DE - t| > \epsilon' t \Pr[ERROR] \le \delta$.

► Theorem

Algorithm 1 returns an estimate of DE within $(1 \pm \epsilon)$ with probability $\geq (1 - \delta)$.

AMS Sketch for Counting Distinct Element

- Uses pair-wise independent hash function
- Improved space and time complexity
- Worse approximation

Algorithm 4 AMS Counting Distinct Items

```
Initialize
z \leftarrow 0
End Initialize
Process(a_l = (j, \nu))
if zeros(h(j)) > z then
  z \leftarrow zeros(h(j))
end if
End Process
Estimate
return 2^{z+\frac{1}{2}}
End Estimate
```

AMS Sketch for Counting Distinct Element

• Define
$$X_j^r = 1$$
 if $zeros(h(j)) \ge r$ and 0 otherwise. Define $Y^r = \sum_j X_j^r$.

Lemma

$$E[X_j^r] = \frac{1}{2^r} E[Y^r] = \frac{DE}{2^r} Var[Y^r] \le \frac{1}{2^r}$$

Lemma

• Consider the largest level a such that $2^{a+\frac{1}{2}} < \frac{DE}{3}$. $\Pr[z \le a] < \frac{\sqrt{2}}{3}$.

 Consider the smallest level b such that 2^{b+¹/₂} > 3DE. Pr[z ≥ b] < √2/₃.
 Pr[DE/3 < 2^{z+¹/₂} < 3DE] ≥ 1 - 2√2/₃.

AMS Sketch for Counting Distinct Element

► Boosting the confidence Median Trick. Keep C log ¹/_δ copies and return the median estimate

Theorem

There exists a randomized algorithm that returns an estimate of DE satisfying $\Pr\left[\frac{DE}{3} < 2^{z+\frac{1}{2}} < 3DE\right] \ge 1 - \delta$ using space $O(\log \frac{1}{\delta} \log n)$